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Empirical Application of the Housing Market No-Arbitrage Condition: Problems, Solutions and a Finnish Case Study

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***Abstract.** The housing market no-arbitrage condition is a theoretically sound basis to evaluate if housing prices are misaligned. Unfortunately, empirical application of the no-arbitrage condition has notable complications. This article reviews these complications and suggests some solutions to them. In particular, the use of implied expected appreciation derived from the no-arbitrage condition is recommended. Furthermore, the paper shows that the maintenance costs as a fraction of the housing price level substantially vary in time and location, which may significantly affect the equilibrium housing price level relative to rental prices. An empirical application of the no-arbitrage relation using data from ten cities in Finland shows that the housing price level has not been based on high expected housing appreciation during the 2000s.*

***JEL Classification:** R21*

***Keywords:** housing, price, rent, no-arbitrage condition, user cost, bubble*

1 Introduction

Housing prices soared in most of the industrialized countries during the decade preceding the sub-prime crisis. The rapid housing price increase raised questions about the sustainability of the prevailing housing price levels. It is a common view that in a number of countries the recent substantial housing price decline has not been only due to the drastic deterioration in the economic fundamentals but also because of a notable overvaluation of housing prior to the crisis. Regardless of the magnitude of the drop in the housing price level, the question regarding potential misalignments in housing prices remains: have prices fallen enough or too much or is housing still overpriced?

Given the important role that housing wealth seems to play on aggregate demand through its impact on construction activity, household consumption and on the credit market, it is worthwhile to assess reliably the prevailing housing price level in order to evaluate better the prospects of the economy. Consequently, different institutions and scholars have put a lot of effort to evaluate if housing

prices misaligned in a number of countries and cities. Various methodologies from econometric analyses to simple ratios between housing prices and some fundamentals have been employed in the sustainability analyses. Unfortunately, the evaluation of housing price level is anything but an easy task and all of the assessment methodologies have their complications.

A theoretically sound basis for evaluating if housing prices are misaligned is given by the so-called housing market no-arbitrage condition introduced by Poterba (1984). The no-arbitrage condition states that in equilibrium the user cost of housing should equal rental price level of a similar dwelling. In addition to its appealing intuition to the housing market, the advantage of the no-arbitrage condition is that it takes the future expectations into account by a single expected housing appreciation term. Recently, the no-arbitrage condition has been utilized in a number of studies (e.g. Finicelli, 2007; Girouard et al., 2006; Himmelberg et al., 2005; MacCarthy and Peach, 2004) to assess if housing prices are misaligned in one or more countries or cities. Unfortunately, the empirical application of the condition involves several complications and all of the above mentioned examinations have their problems. Nevertheless, careful analysis and data selection may enable a more reliable evaluation of housing price level based on the no-arbitrage relation than based on the alternative methods.

This article discusses the complications of utilizing the no-arbitrage condition in an empirical analysis. Furthermore, a case study employing the no-arbitrage condition using data from a number of cities in Finland is conducted. In connection with the discussion and the case study, some solutions to the presented problems are suggested.

It is proposed that the implied expected future appreciation should be reported when evaluating the prevailing housing price level through the no-arbitrage relation. The paper further suggests that it is more convenient to employ real instead of nominal appreciation rates in the analysis. Moreover, it is argued that the variation in maintenance costs should be catered for to get reliable results, since the fraction of maintenance costs of housing prices varies in time and region. Furthermore, if the price-to-rent ratio is employed, due to simplicity reasons for instance, the rental price net of maintenance costs should be used instead of the traditionally employed “gross” price-to-rent ratio. In growing regions the gross price-to-rent ratio is likely to overstate current housing price level. It is also noted that the differences in maintenance costs weaken the comparability of the conventional (gross) price-to-rent figures across cities.

According to the empirical case study, housing price level has not been based on high expected appreciation in the Finnish cities in the 2000s. This finding lowers the fears for a price bubble and thereby for a need for a drastic housing price drop in Finland. In addition, empirical analysis suggests that either the Finnish households’ risk premium on owner-occupied housing is substantially greater than the risk premium figures suggested in the previous literature or Finnish households have constantly underestimated the future housing appreciation since the early 1990s. A partial explanation for the constantly smaller user cost than rental cost may also be exhibited by the credit constraints.

Previously, Girouard et al. (2006) have applied the no-arbitrage relation to Finnish data. In line with the results in this paper, the analysis by Girouard et al. implies that in 2004 housing price level was somewhat lower than its fundamental level in Finland. Girouard et al., however, use national level data and the analysis has also some other complications. Since there may be notable regional differences in the user cost-to-rental price ratio within a country, it is reasonable to conduct the examination to distinct regional housing markets as in this article.

The outline of the paper is as follows. The next section delineates the price-to-rent and price-to-income ratios in the case areas and ponders the problems of relying on such ratios. In the third section the no-arbitrage condition is presented and complications with its empirical application are discussed. The empirical case study is conducted in section four. In the end, summary and conclusions are drawn.

2 Price-to-income and price-to-rent ratios

Because of their simplicity and appealing intuition to the general public, housing price-to-rent ratio and housing price-to-income ratio have repeatedly been employed e.g. by credit institutions and by the media to justify views concerning the sustainability of existing housing price level. However, the use of either of these simple ratios has considerable problems and may lead to flawed conclusions, as discussed below.

Fig. 1 exhibits the price-to-income ratio (P/Y), i.e. the average housing price level per square meter divided by the annual disposable income per capita, in ten Finnish cities. Fig. 2, in turn, pictures the price-to-rent ratio (P/E), which equals the average housing price level per square meter divided by the average annual rent per square meter.¹ The dash lines in the Figures show the average level of the ratios during 1985–2008. Two different price-to-rent ratios are shown in Fig. 2. “Gross P/E ” corresponds to the ratio that is typically presented in various reports, i.e. the ratio where maintenance costs are not deducted from the rental cash flow. A more justifiable ratio, however, is the ratio between housing prices and rental cash flows net of maintenance costs. This ratio is called “net P/E ” here.

In each of the ten cities, P/E values have declined from 1985 to 2008 despite the rapid increase in housing prices during the sample period.² That is, rental prices have grown even faster than asset prices. This is partly due to the changes in the Finnish rental market regulation. Lifting in the rent ceilings started in the late 1980s and the rent regulation was finally released in several stages during 1992–95. During the housing price overshoot of the late 1980s, which followed the credit market liberalization, both P/Y and P/E were substantially above their average levels. Then, in the early 1990s, the ratios dropped below their long-run averages. Since then P/Y has climbed up. However, in many cases P/Y has increased only slightly even though real housing prices have risen by at least 58%

¹ The income data are published by Statistics Finland. The data on housing prices and rents are described in section 3.5 and on maintenance costs in section 3.2.

² Nominal price level rose between 105% (Rovaniemi) and 229% (Tampere) over 1985–2008. The real appreciation varied between 20% and 94%.

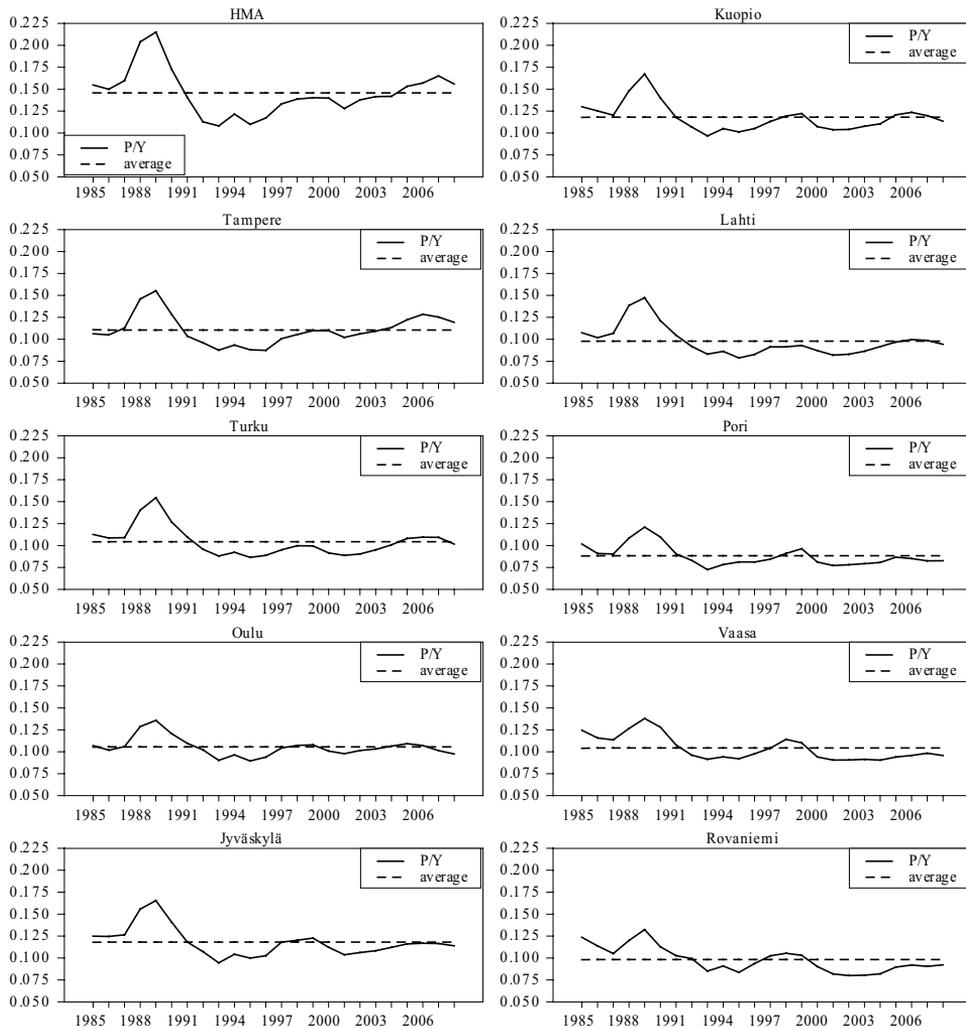


Fig. 1. Price-to-income ratios.

in all of the cities from 1995 to 2008. That is, in many cities the income growth has almost equalled housing appreciation after the severe recession in the mid 1990s. Moreover, with the exception of HMA and Tampere, P/Y was lower in 2008 than in 1985 and than the average ratio during 1985–2008. In 2008, the average price of one square meter equalled approximately 16% of the average annual disposable income per capita in HMA. The corresponding value in Tampere was 12%. The net price-to-rent ratio, in turn, was slightly below its long-term average even in HMA and close to its average in Tampere. In the other eight cities the price-to-income ratios in 2008, varying from 8% to 11%, were below the estimated averages.³

In 2008, the net P/E was at a lower level relative to its long-run average than the gross P/E in each of the cities. This is because the rental price level has

³ In the case of Vaasa (1989–2008) and Rovaniemi (1991–2008) the average values of the price-to-rent ratios are based on even shorter sample periods than in the other cases.

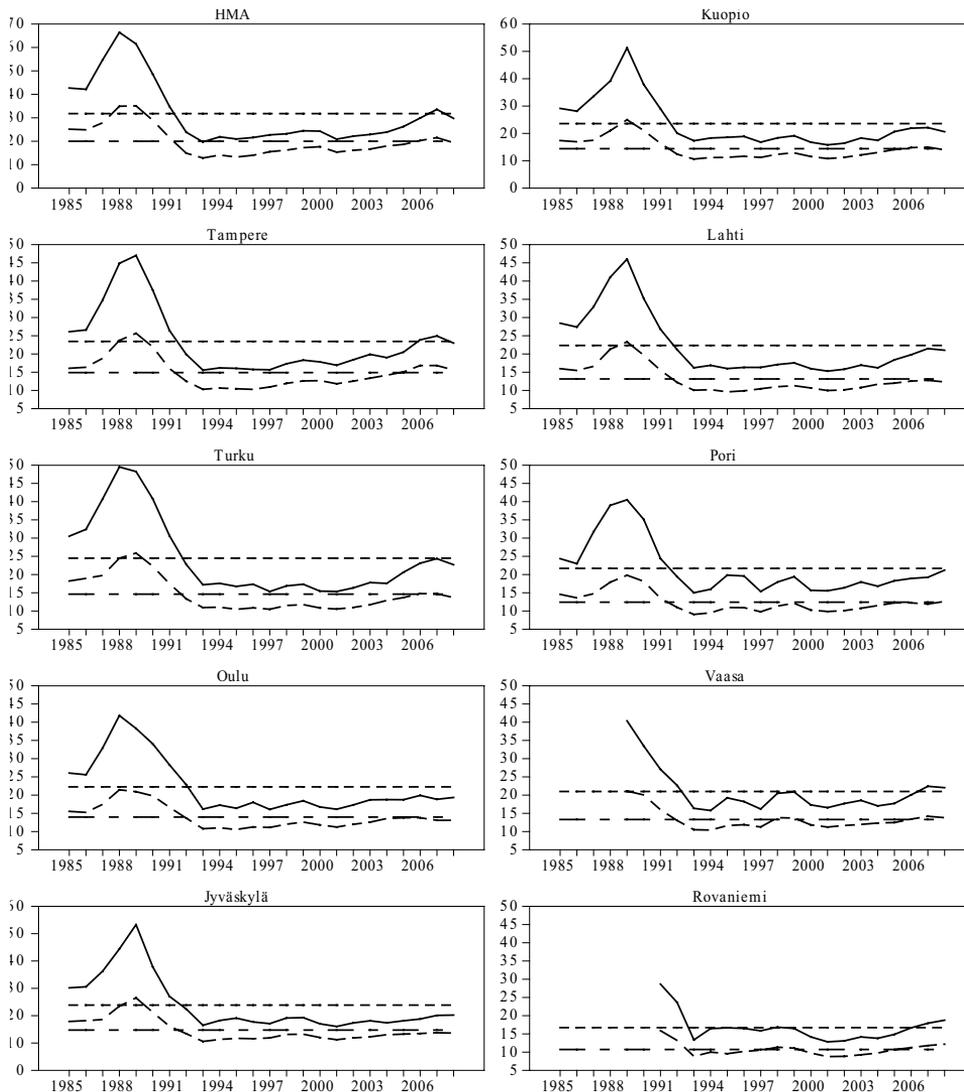


Fig. 2. Price-to-rent ratios. The continuous curve shows the net P/E and the dash curve presents the gross P/E. The linear lines show the average P/E values during 1985–2008.

increased notably faster than the maintenance costs. That is, the conventionally used gross P/E ratio is likely to overstate the current housing price level relative to fundamentals in growing areas, i.e. in areas where rental price growth is relatively rapid. The impact of relative differences between rental prices and maintenance costs is discussed in more detail in section 3.2.

POINT 1: If P/E is used to evaluate housing price level, the *net* P/E should be employed instead of the conventionally used *gross* P/E, since the maintenance costs and rental prices do not, in general, grow at the same rate.

P/Y works as a kind of “affordability index” of housing. However, there is no particular reason to assume that P/Y should be constant over time. In fact, depending on the elasticities of supply of labor and of supply of housing and on the driving forces behind metropolitan growth, P/Y can be decreasing, constant or growing in time (see DiPasquale and Wheaton, 1996, 155–165). In addition, Meen (1996) shows that a range of policy shocks, changes in the tax regime for instance, is capable of shifting the relationship between housing prices and income. In the Finnish case, especially the credit market liberalization may have significantly influenced the ratio. Hence, it is problematic to use P/Y to evaluate the prevailing housing price level.

Theoretically, the use of P/E is somewhat more justifiable. There are reasons to assume that the housing price and rental price levels are tightly linked. For investors, rent represents the incoming cash flow from the housing investment. For owner-occupants, in turn, the rental price level exhibits the cost they would face if they did not own the dwelling. In other words, rental price is an implicit positive cash flow for the owner-occupants. However, there are factors, such as the interest rates and expectations concerning income and population growth in the area, which vary over time and across regions and may affect the ratio significantly. Changes in some of these factors are likely to be highly persistent. Therefore, the equilibrium P/E may well notably alter over time and in many cases is not likely to be stationary. Indeed, none of the P/Y ratios and only a couple of the P/E ratios appear to be stationary (see Table A1 in the Appendix). Naturally, shortness of the sample period may somewhat affect the unit root test results.

Finally, even if P/Y and P/E were constant over the long horizon and even if there had not been significant institutional alterations, it would be problematic to compare current ratios to the averages during 1985–2008. This is because it is unclear whether the sample represents a “normal” time period, i.e. whether the average ratios during the sample have equalled the long-term equilibrium ratios.

POINT 2: It may be misleading to base the assessment of prevailing housing price level on the comparison of the price-to-income and price-to-rent ratios with their long-term averages.

The price-to-rent ratio has also been utilized to compare the housing price level between different areas. While the net P/E was 30 in HMA in 2008, the ratio was less than 20 in Rovaniemi. Does this necessarily mean that housing prices were higher relative to fundamentals in HMA than in Pori? No it does not. As will be explained in the next section, equilibrium P/E is influenced by factors such as risk premium, expected growth rate and the share of the physical structure of the total housing price level that may substantially differ between cities. Hence, the comparison of the P/E values between cities does necessarily not tell us anything about potential misalignments in regional housing prices. Someone might claim though that the 3.4% net rental yield in HMA is all too low for housing to be correctly priced in the area. It is not that obvious, however: if real housing price level stays constant, the total return on housing in HMA is actually 3.4% in *real* terms. In fact, catering for the average real appreciation rate of 2.3% (which

would lead to 5.7% real total return) in the HMA over 1985–2008, the relatively low rental yield does necessarily not sound all that bad.

POINT 3: Comparison of the P/E values between different regions cannot, in general, be used to evaluate the "fairness" of regional housing price levels.

3 No-arbitrage condition

A theoretically sound basis for the evaluation of the prevailing housing price level is given by the housing market no-arbitrage condition (NAC). The theoretical foundation of NAC lies on the asset market approach of housing markets introduced by Poterba (1984).⁴ The no-arbitrage condition states that the user cost of owner-occupied house (U) should equal the rent (P^R) of a similar dwelling. More specifically:

$$U_t = P_t [\sigma_t (1 - T^m) r_t^m + (1 - \sigma_t)(1 - T^f) r_t^f + \lambda_t + M_t - i_t] = P^R_t \quad (1)$$

where M_t denotes the maintenance costs of housing as a fraction of house price and i_t stands for the expected nominal rate of future housing appreciation. M_t includes property taxes and depreciation. Here, depreciation refers to the maintenance and repair costs that are necessary to maintain constant quality of the structure. The rest of the term in the parenthesis expresses the opportunity cost of capital. In general, only a fraction σ of the value of a house is financed by mortgage. The cost of the mortgage capital equals the after-tax mortgage rate, $(1 - T^m) r^m$, where T^m is the tax deductibility of mortgage interest payments in taxation and r^m is the before-tax mortgage rate. The opportunity cost of housing capital is completed by the after-tax risk-free interest rate $[(1 - T^f) r^f]$, where T^f is the tax rate for capital income] on the part of the dwelling financed by equity plus the additional risk premium (λ) to compensate homeowners for the higher risk of owning vs. renting.⁵ Note that all the variables in (1) are in nominal terms. An inflation correction is carried out in the model by the inclusion of the expected *nominal* housing appreciation.

NAC can be used to illustrate the main problem with the price-to-rent ratio. From the equivalence condition in (1) we get:

$$P/P^R_t = 1/[\sigma_t (1 - T^m) r_t^m + (1 - \sigma_t)(1 - T^f) r_t^f + \lambda_t + M_t - i_t] = P/E^{eq}_t \quad (2)$$

⁴ Poterba (1984) focuses on the price of housing structures only. Nevertheless, the same basic idea applies to housing prices consisting of both the structure and the site.

⁵ Englund et al. (1995) incorporate also a term catering for the effect of housing ownership on the wealth tax. In Finland, owner-occupied housing is not taxed in the wealth taxation. Hence, assuming that if the capital was not invested in housing, it would be invested in an investment form that is taxed in the wealth taxation, the wealth tax reduces the user cost of owner-occupied housing (there is an extra tax benefit). However, the rules regarding the wealth tax are complicated in Finland, and it would be highly problematic to try to include a wealth tax benefit term in (1) in an empirical application.

It is evident that the “equilibrium” P/E, denoted by P/E^{eq} , is the larger the lower is the risk-free opportunity cost of capital and the greater is the expected future housing price growth. Furthermore, the risk premium and maintenance costs affect P/E^{eq} . The discount rate in the parenthesis in (2) cannot generally be assumed to be stable over time or stationary. Hence, one cannot make reliable conclusions on whether housing is fairly priced or not solely by comparing the P/E ratio with its long-term average value. Note also that it is clear from (1) and (2) that the sensitivity of housing prices to changes in the fundamentals, i.e. in the factors that affect the rental price level or variables in the discount factor, is higher when the discount factor is low.

POINT 4: The sensitivity of housing prices to changes in fundamentals is greater in regions where and periods when the discount factor is low.

Although NAC is typically used to evaluate housing price level from the point of view of owner-occupied housing, the idea also applies for investment housing. From an owner-occupant’s viewpoint the intuition behind NAC is that the cost of same good (housing service) should be equal whether one owns the dwelling or not. From an investor’s point of view, in turn, NAC indicates that the expected return on housing equals the required return. If expected return was higher than the required rate, investors would bid up prices back to the equilibrium.

In Finland, the capital gains tax is imposed on the rental returns (net of maintenance costs and interest payments on the loan that is borrowed to buy the dwelling) whereas the owner-occupants’ imputed rental income is not taxed. Moreover, housing appreciation is taxed when a rental dwelling is sold. Capital gains on owner-occupied housing, instead, are generally tax-free. Hence, because of the tax benefits of owner-occupied housing, housing is worth less for investors than for owner-occupants if the other variables in (1) are the same for both groups. Therefore, if housing is of equal worth to owner-occupants and portfolio investors, the required return set by investors has to be lower or the future expectations more positive than those of the owner-occupants. Anyhow, as most of the privately financed dwellings are owner-occupied in Finland as well as in most of the other developed countries, the owner-occupants’ view is taken in the forthcoming analysis.

Unfortunately, there are several complications with the empirical application of NAC. First, due to the high transaction costs and low liquidity of housing as well as due to households’ liquidity constraints, in reality, there can be slight divergence from the presented relation even if the market participants are fully rational. In particular, if households are tightly liquidity constrained, user cost may be lower than the rental cost for sustainable periods.⁶ If liquidity constraints are not significant, instead, the relation should hold in the long run. The transaction costs

⁶ Meen (1996) adds a shadow price of the credit rationing constraint in the user cost formula to cater for the credit constraints. Unfortunately, such a shadow price measure is not available in most countries, including Finland.

are far from straightforward to take account of.⁷ Therefore, transaction costs are usually ignored in the empirical analyses employing NAC as well as in the empirical analysis in this paper. A related problem is the measurement of an appropriate investment horizon. After all, due to the transaction costs and relatively low liquidity of housing it is reasonable to assume that the typical planned investment horizon of housing is relatively long. As DiPasquale and Wheaton (1994) state, “*The expected price term refers to current or next period price inflation only if there are no transaction costs to altering housing consumption. When transaction costs impede mobility, the price term must consider planned holding periods*”.

Second, the measurement of risk premium and of expected appreciation is difficult. Specifically, how should one estimate these two variables and are these variables constant through time? Also the measurement of the other variables may include some complications. The problems are similar to the other methods to assess housing prices and, as in any empirical analysis, incorrect measurement of the variables may give rise to misleading conclusions.

Notice that the discount factor may differ between households. In this paper the attractiveness of owner-occupation is assessed from the viewpoint of both existing owner-occupants and of potential first-time home-buyers. In practice, the separation is of importance only concerning the values of σ , T^m and possibly λ . It is worthwhile to analyse the potential first-time buyers' view separately, since Ortalo-Magné and Rady (2006) show that the importance of first-time buyers on housing market dynamics is likely to be particularly important, at least in the relatively short horizon. Despite the complications, careful analysis and data selection enables one to make relatively reliable conclusions based on NAC, and the relation between user cost and rental prices has been utilized in a number of papers studying housing price dynamics or examining the fairness of housing prices. Below, the complications regarding the empirical application of NAC are discussed in more detail and past empirical studies are reviewed. Furthermore, some potential solutions to the exhibited problems are suggested. In connection with the discussion, the data employed in the empirical case study are delineated.

3.1 Expected appreciation

Typically, simple and somewhat arbitrary assumptions concerning the expected appreciation are used in the literature. Girouard et al. (2006) and Poterba (1992) assume that the expected housing price increase equals the expected rate of overall inflation. Poterba approximates the expected inflation as the arithmetic average and Girouard et al. as the moving average of inflation rate in the five preceding years. Englund et al. (1995), in turn, employ the average annual housing appreciation during the 1980s, while Finicelli (2007) proxies the expected capital gain as the sum of long-term inflation expectations and historical growth in real rents. Himmelberg et al. (2005), who include also a forward-looking component by adding the spread between long- and short-term interest rates in the user cost

⁷ Transaction costs include brokerage fees, taxes, paperwork, information gathering as well as the financial and psychological costs of moving.

formula, use the average real growth rate of housing prices from 1940 to 2000 as a proxy for expected real appreciation. Furthermore, Smith and Smith (2006) use the same arbitrary expected appreciation values (in the base case 3%) for a number of different areas in the US, even though the actual expectations are likely to substantially differ between different regions.⁸ DiPasquale and Wheaton (1994), instead, calculate, what they call, rational expectations. Even these expectations are based solely on historic values of the fundamental variables.

The measurement of expected appreciation is of major importance, since different methodologies can lead to different conclusion about the extent of misalignment in housing prices. Obviously, the use of purely backward-looking expectations utilized in several papers may induce misleading conclusions. Past housing appreciation or overall inflation do necessarily not represent well the rational expectations. For example, if housing prices have risen rapidly during the past few years and are currently notably above their fundamental level, backward-looking expectations based on relatively short history imply fast housing inflation also in the future. On the contrary, rational agents should take the prevailing overpricing into account. That is, rational agents would cater for the adjustment of housing prices towards the fundamental level and, therefore, the forward-looking expectations would predict a substantially lower appreciation figure.

Nevertheless, because it is extremely difficult to evaluate correctly the rational expectations at a given point of time and because based on the empirical literature the expectations appear to be, to some extent, backward-looking in the housing market, it is understandable that mainly backward-looking expectations have been employed in the user cost literature. In particular, the longer the horizon is, the harder it is to predict the development of the price level. The expected appreciation figure caters for the expectations concerning income and population growth (and thereby rental growth) in the area as well as for expected interest rate movements. Therefore, assuming that the expectations are not notably different from the historical population and income growth and that nothing radical is expected to happen in the real interest rate level, it is reasonable to employ the average appreciation during a long period in the past.⁹ If historical averages are utilized, i should be based on the average real, not nominal, appreciation. This is because inflation figures typically were substantially higher in the past than today. Nevertheless, a number of studies employ the average nominal appreciation figure.

POINT 5: If a historical average housing price growth rate is used to estimate the expected appreciation, the real, instead of nominal, growth rate should be used.

In the empirical analysis below, three different appreciation assumptions are

⁸ Smith and Smith use the present value approach in their analysis. The present value approach is basically equivalent to the no-arbitrage approach.

⁹ “Long” period refers to a time span of preferably several decades, or at least a period which includes several economic cycles. The use of an average appreciation that is based on a couple of preceding years only may cause significant problems, as discussed above.

employed: the first two assume a constant expected real appreciation at zero and at the average rate during 1987–2008¹⁰, whereas the third model presumes “perfect foresight”. The perfect foresight model assumes that households have perfectly foreseen the changes in housing prices at a one year horizon.

Another option is to calculate the “implied” expected appreciation, i.e. the magnitude of i at which the user cost equals rental price level. If the implied growth rate is very high, there are concerns that the current price level is based on overly optimistic and even speculative growth expectations. On the contrary, low implied appreciation would suggest that the housing price level is not based on high price growth expectations and, therefore, the risk of notable overpricing is relatively small. By using the implied expected appreciation it is also easy to see what happens if one of the variables in the discount factor is altered. For instance, if one thinks that the risk premium for housing should be one percentage point higher than used in the calculations, he can just add one percentage point to the presented implied appreciation figure to find the situation based on his own views. Similarly, one can easily get answers to questions such as “what if the interest rate was higher” by employing the implied appreciation rate. The fact that the ratio between user cost and rental cost is quite sensitive to variation in factors such as the level of maintenance costs, risk premium and interest rate makes the use of implied expected appreciation worthwhile.

POINT 6: It is reasonable to report the “implied” expected appreciation when assessing housing prices by the no-arbitrage condition.

3.2 Maintenance costs

The maintenance costs are typically assumed to be a constant fraction of housing prices. In Englund et al. (1995), Finicelli (2007) and Poterba (1991), M equals 4%, 5.5% and 4%, respectively, whereas Girouard et al. (2006) set $M+\lambda$ to a constant 4%. Furthermore, even though M is likely to vary substantially between different areas, Himmelberg et al. (2005) assume that M equals 2.5% in all the cities included in their analysis.

Generally it is not reasonable to assume that the maintenance costs are a constant share of the value of housing. Typically, in a growing metropolitan area the appreciation of land accounts for a significant part of the housing price growth.¹¹ As it is, in general, only the structure that depreciates, it is likely that M decreases in the long run in a growing metro area. That is, the evolution of M alone can cause the gross P/E to trend upwards in the long run. Furthermore, it is

¹⁰ The average growth rates are calculated based on the hedonic housing price indices provided by statistics Finland.

¹¹ According to Rosenthal and Helsley (1994), Davis and Heathcote (2007) and Davis and Palumbo (2008) the value of land can account for over half of the price of housing in large cities. In fact, Davis and Heathcote and Davis and Palumbo find that land’s share of aggregate home value has been trending upwards in the U.S. Bostic et al. (2007) show that even in a relatively small city with no notable geographical restrictions the value of land can grow substantially faster than the value of the structure.

expected that in cities with higher housing prices, i.e. higher value of land, M is smaller than in the regions with lower housing price level. In other words, other things being equal, the gross P/E is expected to be larger in bigger cities. Malpezzi (1999) finds this to be true in the US market.

In this study, proxy for M is based on the average per square meter maintenance costs (including the repairs needed to maintain constant quality) of privately financed flats reported by Statistics Finland. Except for HMA, there are no statistics at the city level. Hence, the reported values for corresponding greater geographical area are used for the other nine cities. Major part of the flats in each of the greater regions is located in the case cities, which limits this data problem. The maintenance costs are reported annually. Hence, departing from the previous literature, M is allowed to vary in time.

Expectedly, M varies substantially over time and between the case cities. The smallest figure is 1.2% in HMA in the peak of the price overshoot in 1989, while the largest M is 4.4% in Rovaniemi in 1995. In 2008, M was smaller in all the cities (except for Lahti) than the average during 1989–2008, varying from 1.8% in HMA to 3.2% in Pori. This is not surprising, since housing prices have grown fast during the last decade while construction costs have increased only slightly, i.e. the share of land value of the total price of housing has been trending upwards. In general, M is smaller in larger cities, i.e. cities with higher housing prices, as expected.

POINT 7: The fraction of maintenance costs of housing prices varies over time and between regions. This should be taken into account in the calculations to get reliable results. Furthermore, the differences in maintenance costs weaken the comparability of the conventional (gross) P/E figures between cities.

3.3 Risk-free opportunity cost of capital

In the literature, usually the after-tax mortgage interest rate is assumed to equal the risk-free opportunity cost of capital tied in owner-occupied housing (see e.g. Englund et al., 1995; Girouard et al., 2006; McCarthy and Peach, 2004; Poterba, 1991). At the same time, the ratio of the mortgage debt to the value of the owner-occupied house is assumed to be 100%. An exception is the analysis by Himmelberg et al (2005), where the risk-free opportunity cost of capital is measured as the ten-year US Treasury rate. Himmelberg et al. use the mortgage rate only to calculate the tax benefit. Also they employ a 100% debt-to-value ratio, however.

In reality, the debt-to-value ratio of most owner-occupiers is substantially below one. Even the first-time home-buyers usually need a down-payment. If the risk-free interest rate equals the mortgage rate and T^m equals T^f , the assumption regarding the debt-to-value ratio does not matter. However, if $r^m \neq r^f$ or $T^m \neq T^f$ the debt-to-value ratio is of significance. In general, banks include a risk premium to their lending rates. Hence, it is reasonable to assume that in most periods $r^m > r^f$ at least in Finland, where the interest rate of most housing loans is fixed only for 12 months at a time. Therefore, in the Finnish case the use of 100% ratio is likely

to exaggerate the user cost somewhat, since T^f has equaled or been close to T^m . Since the debt-to-value ratio is, in general, substantially greater for the first-time buyers than for the households that already are owner-occupants, the distortion is emphasized if the housing price level is evaluated from the viewpoint of the owner-occupants. In Finland, for instance, the estimated average mortgage-to-value ratio was approximately 25.4% in 2008.¹² In the first-time home-buyer case, in turn, a 90% ratio is employed in this study.

POINT 8: Depending on the difference between risk-free interest rate and mortgage rate and on the tax rules, the employed debt-to-value ratio may notably affect the results.

Another problem is whether to use the average interest rate on the new mortgage withdrawals or on the whole outstanding mortgage stock. Evidently, in the calculations concerning potential first-time buyers it is sensible to use the rate on new withdrawals. Concerning the owner-occupants, on the contrary, the average rate is a more prominent stand point. Anyhow, the spread between the two rates has been, in general, only slight in Finland. The spread between new contracts and old stock was quite large, from .71% to 1.71% during 1989–1992, however. Due to a change in tax rules in 1993, T^m equals the average marginal income tax rate prior to 1993 and the capital income tax rate or 30% (first-time buyers) from 1993 onwards.¹³ Furthermore, it is assumed that the deduction ceiling is not binding in either of the groups. In the Finnish case T^f equals T^m with the exception of the 30% rate for the first-time buyers. Finally, the 12 month euribor is used as the risk-free interest rate.

3.4 Risk premium

Also the risk premium is assumed to be constant in the empirical applications. Finicelli (2007), Flamin and Yamashita (2002) and Himmelberg et al. (2005) assume a risk premium of 2%. Poterba (1991), instead, uses a 4% risk premium, while Englund et al. (1995), McCarthy and Peach (2004) and Quigley and Raphael (2004) do not include a separate risk premium at all. Himmelberg et al. claim that even the 2% risk premium may be overly high. Indeed, there are factors that may decrease λ . First, households might derive extra utility from owning a house (e.g. ability to customize the interior or pride of ownership).¹⁴ Second, also renters confront uncertainty. Typically, tenants can expect to have to move more

¹² This figure is likely to somewhat overstate the actual average mortgage-to-value ratio of owner-occupants since the mortgage loan stock includes mortgages withdrawn by landlords to finance their housing investments. Unfortunately, there is no sufficient data to estimate the loan-to-value ratio for the cities separately. Hence, the national time series is employed in the calculations. The housing loan stock statistics are reported by the Bank of Finland.

¹³ The capital income tax rate has varied between 25% and 29%. A great number of the existing owner-occupants can utilize the higher T , however, because they have withdrawn their outstanding mortgages to buy their first home.

¹⁴ This claim is supported by the recent findings of Diaz-Serrano (2009) according to which renters who become homeowners experience a significant increase in housing satisfaction even if the characteristics of the owner-occupied dwelling are identical to those of the rental dwelling.

frequently than owner-occupants, and the future development of rental prices is uncertain. In fact, owner-occupation may work as a hedge against the risk of unanticipated future rental price movements (see Sinai and Souleles, 2005). On the other hand, renting enables better diversification of investment portfolio for a typical household. Note also that the risk premium should be the greater the higher the debt-to-value ratio is, since the use of leverage increases the volatility of the return on equity.

The assumption of constant risk premium may involve similar problem to the constant M . Himmelberg et al. (2005) note that the risk premium is likely to be larger in cities with high housing prices. The rationale behind this is that in cities with high housing prices the value of developed land is generally high relative to the construction costs. As the value of land is generally substantially more volatile than the construction costs (see e.g. Davis and Heathcote 2005; Somerville, 1999), housing prices are likely to be more volatile in cities with high housing (i.e. land) prices.

The higher volatility in larger cities can be rationalized also through the NAC framework. Other things being equal, smaller M leads to smaller discount factor and thereby to a greater sensitivity of housing price to changes in fundamentals. Therefore, since in a growing metro area the sensitivity of housing prices with respect to fundamentals is likely to increase over time, the risk premium may trend upwards in time. On the other hand, in many cases the riskiness of larger cities is diminished by the typically wider economic base compared to smaller cities: risk related to growth of different industries is likely to be better diversified in larger cities. Hence, the increasing trend in the risk premium may be offset if the industrial mix of the city widens as the city grows.

A perfect model would allow the risk premium to vary over time. Because it is extremely difficult to estimate time-varying risk premium reliably, a constant risk premium is assumed also in this study. However, due to the rental market liberalization, λ is assumed to be greater prior to 1993 than currently. Following Flamin and Yamashita (2002), in the owner-occupant case the post 1992 λ is set to 2%, whereas the pre 1993 λ equals 3%. As the risk premium is likely to be the larger the greater the debt-to-value ratio is, the corresponding risk premiums for the potential buyers are 2.5% and 3.5%.

Obviously, the risk premiums employed in this study as well as in the previous literature are somewhat arbitrary. During 1987–2008 the return on Finnish housing over the 12 month euribor was on average 1.9% per annum. This is in line with the suggestion of a 2% risk premium by Flamin and Yamashita (2002) and implies that the risk premiums employed in this study are of reasonable magnitude.

Anyhow, as the implied expected appreciation rates are reported in the empirical analysis below, it is easy to adjust the results for different risk premium assumptions. One only needs to add the increase or decrease in λ to the reported implied appreciation value. The same, naturally, applies if one thinks that the risk premium should vary between the cities.

3.5 Comparing user cost with rental cost

Due to the high transaction costs and low liquidity of housing as well as because of the liquidity constraints faced by households, user costs may somewhat diverge from rental prices even if the market participants are fully rational. A large divergence, however, would imply an existence of housing price misalignment, assuming that the employed data is sufficiently reliable and the presumptions are realistic. In other words, NAC implies that if the user cost-to-rent ratio, i.e. the annual user cost divided by the annual rental payments, is substantially above (below) one, housing is overvalued (undervalued). In reality, the exact actual ratio cannot be observed and the analysis has to be based on best possible approximations.

In some cases, a direct comparison between the user cost and rental cost is not possible. For instance, Himmelberg et al. (2005) compare the user cost-to-rent ratio with its long-term (25 years) average, since their data does not allow for direct comparison. Given that in the long run U should equal P^R , it may be reasonable to assume that the long-run average, indeed, shows the equilibrium. However, if there are price bubbles or institutional changes during the sample period, the estimated long-run average may be distorted. In particular, changes in credit availability in the form of credit market deregulation and other innovations in the credit market affect credit constraints and may thereby cause a structural change in the user cost-to-rent ratio. Thus, if possible, the user cost should be compared directly to the rental cost to make reliable conclusions.

POINT 9: Comparison of the estimated user cost-to-rental cost ratio to its historical average may result in misleading conclusions. Hence, direct comparison between the costs should be made if possible.

In the literature, the considered investment horizon usually is one year. A year is a considerably shorter period than the typical holding period of housing. However, the one-year horizon does not mean that the house will necessarily be sold after a year. It is merely the planning horizon employed in the calculations. The relatively short horizon makes the calculations simple.

One potential shortcoming with an analysis employing NAC is the comparability of housing price and rental price series. In general, the sales price data are based on different dwellings from the rental price data. In this study, the rental price series represent the average rent per square meter in privately financed rental dwellings.¹⁵ For consistency, also the housing price data utilized is non-quality adjusted. The price data does not separate between the owner-occupied and rental units. The general quality of rental housing may be lower than the average quality of housing transacted in the market. Because of this, the perceived user cost-to-rent ratio may exceed the actual one and housing may appear to be more expensive than it actually is. Moreover, the rental price series represent the whole privately financed stock of rental dwellings, not only flats in multi-storey buildings as the price data does. This should not matter significantly, however, since most of the privately financed rental dwellings are flats in multi-storey buildings in all the

¹⁵ Both rental price and housing price data are published by Statistics Finland.

cities included in the examination. This differs from the US case, where owner-occupied units are mostly of the single-family, detached unit type, whereas most of the rental units are in multiple-unit buildings as noted by Glaeser and Gyourko (2007). Also the other differences between owner-occupied and rental housing, such as the quality of the neighborhood, are substantially smaller in Finland than in the US, which makes the use of NAC more reliable. Furthermore, the use of price data on only flats, together with the use of per square meter prices, dampens the potential variation in the quality, size and other characteristics between rental and owner-occupied dwellings.

One problem in some of the papers presented here (e.g. in DiPasquale and Wheaton, 1994; Girouard et al., 2006; McCarthy and Peach, 2004) is the use of national level data. It is highly likely that at least some of the variables in the discount factor substantially vary between different regions and cities within a country. This is why this study conducts the analysis at a city level.

4 User cost vs. rental cost in Finnish cities

To be able to utilize the constant real appreciation assumptions, the following formula derived from (1), which includes the expected inflation rate and the expected real housing price growth separately, is employed to estimate the user costs:

$$U_t = P_t [\sigma_t (1 - T^m) r_t^m + (1 - \sigma_t) (1 - T^f) r_t^f + \lambda_t + M_t - \pi_t - I_t] = P_t^R. \quad (4)$$

In (4), π refers to the expected inflation rate and I denotes the expected real housing appreciation. It is assumed that the households expect the overall inflation rate to equal the inflation rate during the preceding year, measured by the change in the cost of living index. In the perfect foresight case, however, the households are assumed to have a perfect foresight also regarding future inflation. Following the previous literature, the considered investment horizon is one year. In the Finnish case this makes sense also because the interest rate on housing loans is typically fixed only for one year at a time. Hence the assumption of constant mortgage rate at a one year horizon is reasonable.

Keeping in mind the complications and assumptions discussed above, the user cost-to-rental cost ratios are pictured in Fig. 3 and Fig. 4. The curves cover 1989–2008, since the data on the mortgage rates are not available prior to June 1989. In equilibrium the ratio should equal one, which is indicated by the straight dash line in the graphs.

Even at the zero real appreciation expectation there appear to be notable overvaluation in none of the cities. On the contrary, if real housing price level is expected to stay constant, owner-occupation seems attractive at the one-year horizon compared to living in a rental dwelling in all the cities. Naturally, owner-occupation seems even more attractive at the average appreciation rate of the past 20 years. It should be noted, however, that the general price level grew relatively fast (4.1%) from 2007 to 2008. This figure may over-estimate the actual inflation expectation for the subsequent period.

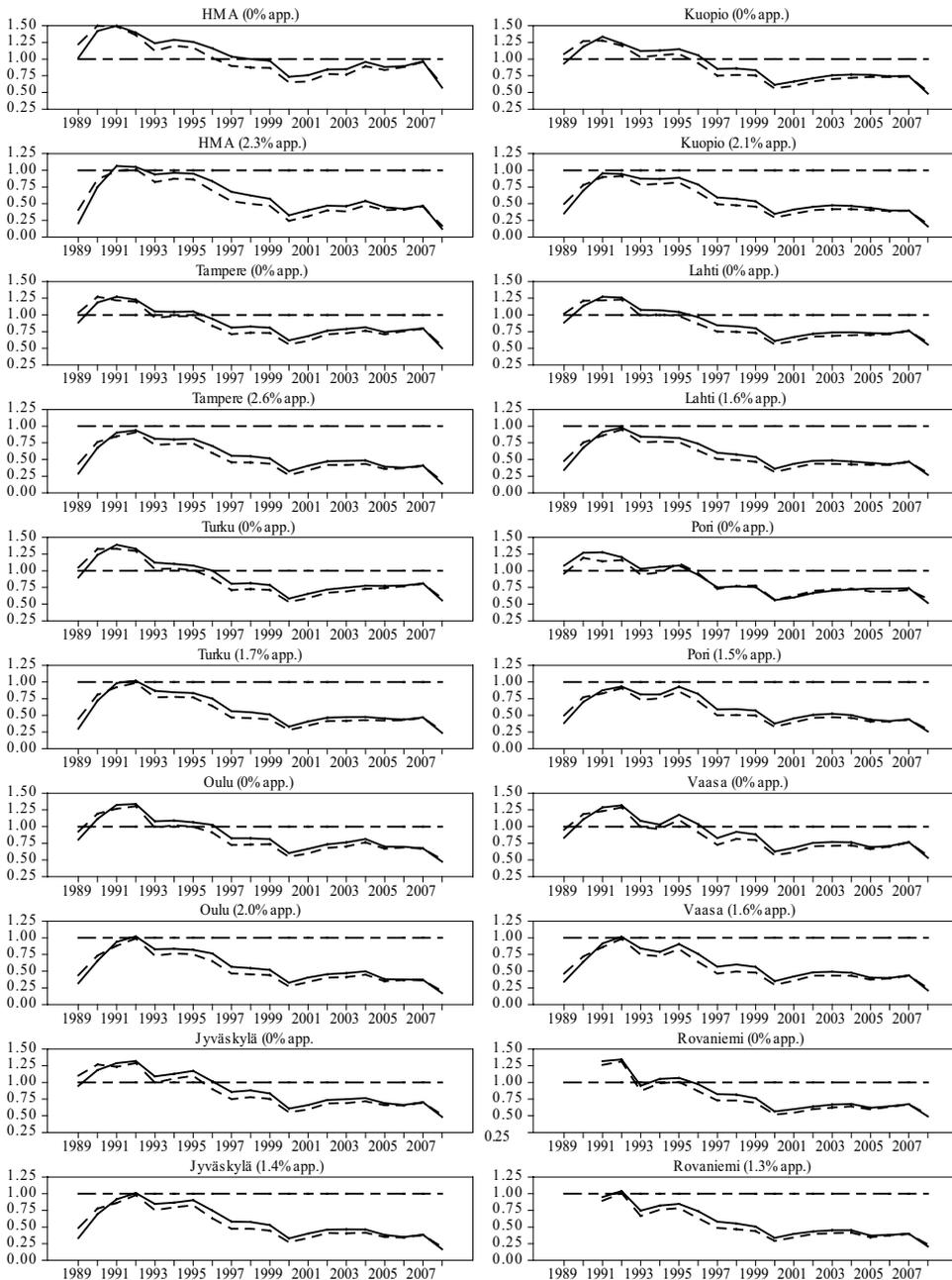


Fig. 3. User cost-to-rent ratios in 1989–2008 in the first-time buyer (continuous curve) and owner-occupant (dashed curve) case based on the constant real appreciation assumption.

In practice, the perfect foresight assumption (Fig. 4) is highly unrealistic. The role of Fig. 4 is to show the user cost-to-rental cost ratios that have actually materialized. The message told by Fig. 4 is that the user cost of owner-occupied flats has been extremely low since 1993, in HMA even negative on average. This

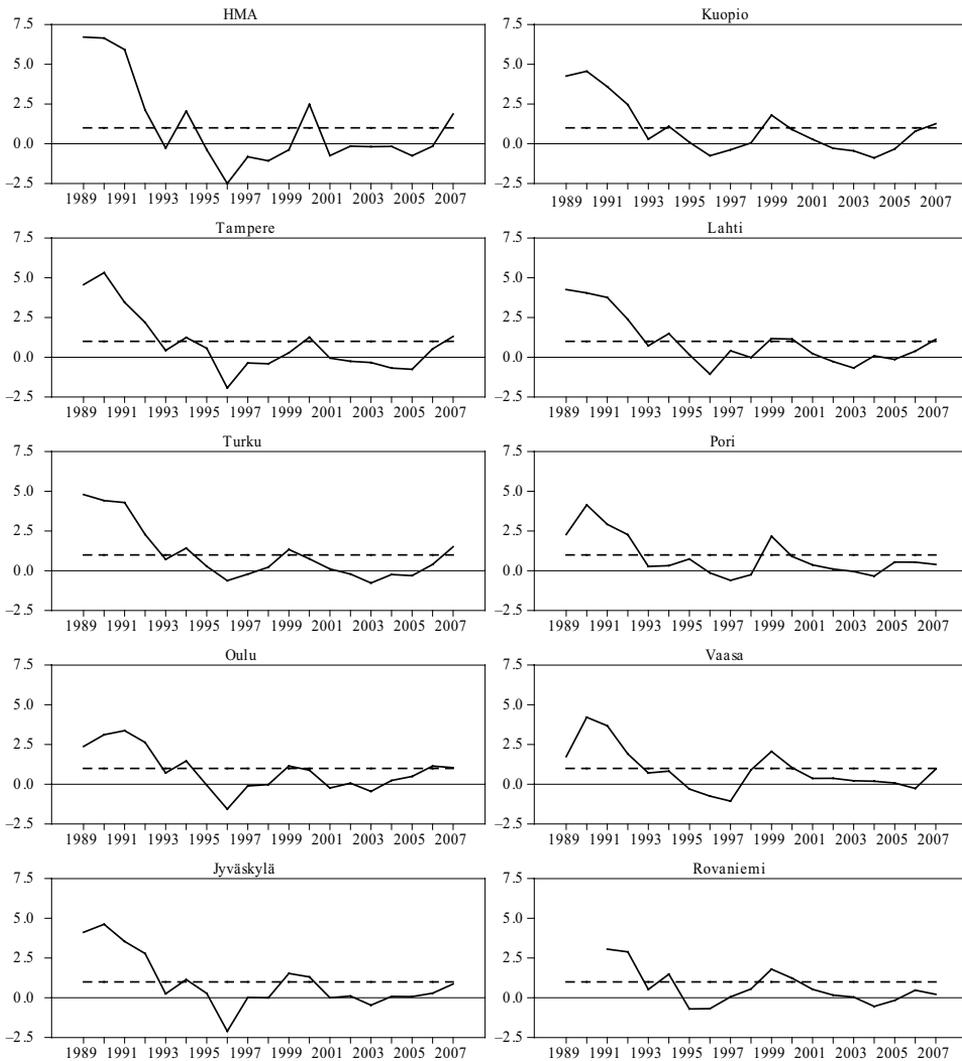


Fig. 4. User cost-to-rent ratios in 1989–2007 based on the perfect foresight assumption.

suggests that, on average, households have highly underestimated the future housing price growth rate after 1993. Another explanation would be that the risk premium for housing is substantially greater in Finland than suggested in the literature. This explanation, however, would require unreasonably large risk premiums. One further factor affecting the divergence may be the credit constraints faced by households. The credit constraints have eased during the last decade, which has probably contributed to the convergence of the ratio towards one.

Even since 1989, a period when housing prices were at an exceptionally high level due to the price overshoot that followed the financial market liberalization, the cumulative rental cost has exceeded the cumulative user cost in each of the cities.

Even if the time value of money is taken into account using a 8.3% discount rate¹⁶, i.e. even if the influence of the low user costs in the later periods is diminished notably, one would have been better off being an owner-occupant compared to living in a rental dwelling in five of the ten cities.

Since the estimated user-cost-to-rent ratios (employing the constant expected appreciation assumption) appear to trend downwards during 1989–2008, it is not surprising that the non-stationarity of the ratios cannot be rejected when (see Table A2 in the Appendix). With longer sample period the existence of unit root in the ratios might be rejected. It has to be noted, however, that due to the loosening in credit constraints there may have been a gradual structural change in the user cost-to-rent ratios. Anyhow, all the unit root test statistics concerning the user-cost-to-rent ratios are closer to the critical values than the statistics regarding the P/E or P/Y series if the same sample period is employed. The values reported in Table A2 cannot be compared to those reported in Table A1 because the sample periods differ.

As discussed above, it is often unrealistic to assume that the expected housing appreciation is constant through time. The perfect foresight assumption in turn, cannot be used to assess current housing prices since future appreciation is not yet known. Furthermore, while the perfect foresight model is clearly forward-looking, it is unrealistic and may lead to false views concerning past housing prices.

Therefore, Fig. 5 pictures the expected real annual appreciation implied by the model (i.e. the value of I_t with which $U_t = P^R$). The calculations indicate that the price level in 2008 was not based on expectations of rapid future housing appreciation. As already shown by Fig. 3, in all the cities the implied real appreciation expectation was negative in 2008. In other words, without changes in the real housing price level owner-occupation appears to be attractive relative to living in rental housing, in general.

The fact that housing price level is not based on high expected appreciation attenuates fears for a notable overpricing of housing. One sign for a housing price bubble would be if prevailing housing prices were founded on overly optimistic expectations regarding future price growth. Obviously, the implied appreciation curves do not support the existence of this kind of speculative bubble in the market. This fact does not change even if a substantially smaller expected inflation rate was used in the computation. For instance, if we assume that the rational expectation of the inflation rate was 1% instead of the 4.1%, all the implied appreciation figures rise by 3.1%. Even in this case the only positive implied appreciation figure would be that in the HMA (.9% in the first-time buyer case). Note further that also the interest rates were at a relatively high level in 2008, and the interest rates have notably declined in 2009.

Furthermore, it has to be noted that potential changes in either the demand side or in the supply side might put downwards pressure on housing prices even though NAC implies that downwards correction is not needed. Other variables

¹⁶ Approximately equal to the average risk-free interest rate during the sample period plus the risk premium of 2%.

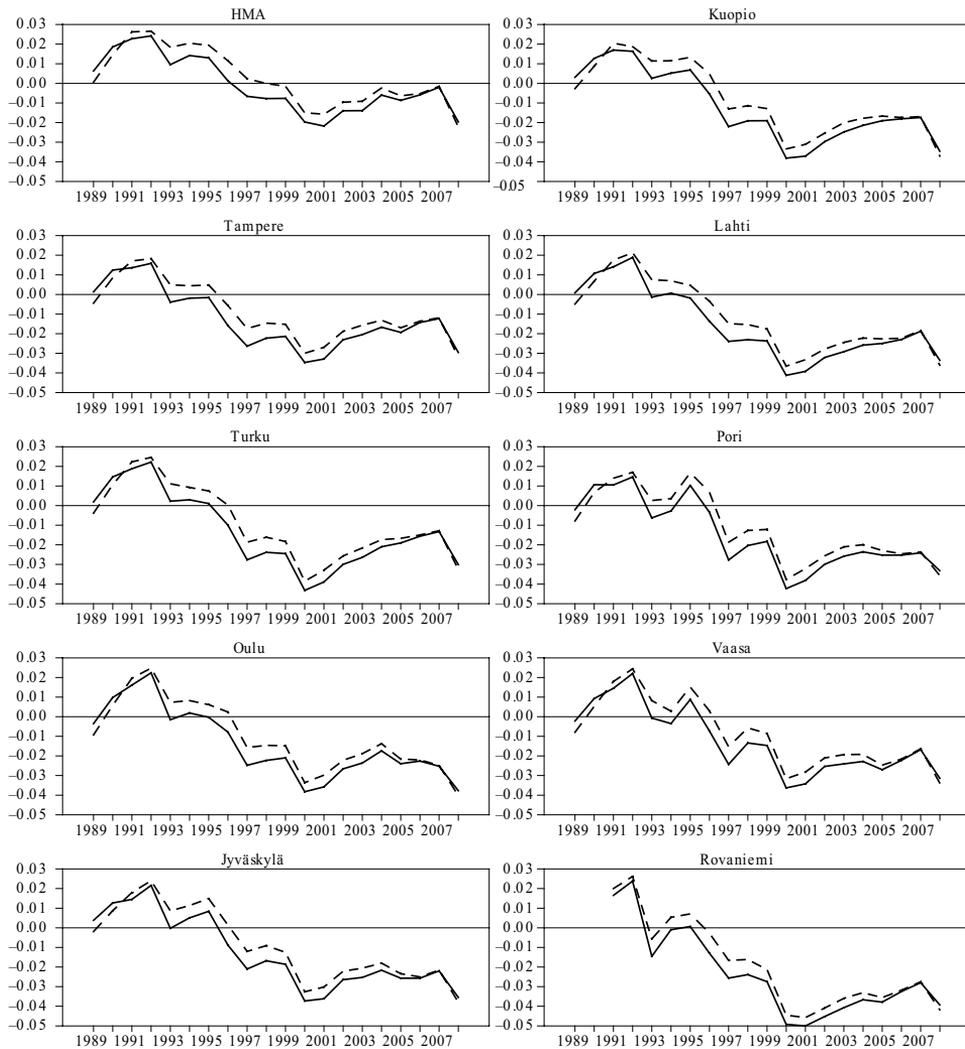


Fig. 5. Expected appreciation implied by the no-arbitrage relation in 1989–2008. The continuous curve corresponds to the owner-occupants' case and the dash curve is for first-time home-buyers.

being constant, greater housing supply would lead to lower prices and rents in the housing market. Moreover, the global financial crisis has already had a negative effect on housing demand.

Surprisingly, the computations suggest that even during the peak of the price overshoot in 1989 the price level did not necessitate huge real appreciation expectations. This may imply that the risk premium is actually somewhat larger than assumed in the analysis. As explained above, one can easily see the implied appreciation expectation assuming higher risk premiums by just adding the extra premium to the implied appreciation rates. Note, however, that a significant factor contributing to the relatively low implied appreciation figures in the late 1980s and early 1990s was the high values of T^m and T^l . The average marginal income

tax rate was almost 53% in 1989. In 1993 the tax rules were altered and T^m and T^r dropped to 25%. It is likely that the high deductibility rates prior to 1993, together with the relatively high inflation and nominal interest rates, fortified the housing price overshoot in the late 1980s as the availability of credit increased remarkably.

Note also that, due to the tax benefit, in the current environment one %-point increase in the nominal (and real) interest rates leads rate causes a from .72% to .70% increase in the implied appreciation rate.¹⁷ However, one %-point increase in the real interest rates due to one %-point decrease in π leads to a %-point decline in I .

The first-time buyers' curves have been at a higher level than the owner-occupants' curves after 1991, since the mortgage rate on new contracts became permanently greater than the risk-free interest rate in 1991. In most cases the implied appreciation was relatively high still after housing prices had substantially dropped in the early 1990s. The figures started to notably decline in the mid 1990s. This was due to the rapid growth of rental prices, decline in M and decrease in the interest rates. Owing to the considerably faster growth of housing prices than rental prices, there was an upward trend in the curves in 2001–2007.

There does not appear to be notable correlation between the implied expected appreciation and the next period actual housing price growth. In fact, over the whole sample period the correlation between the implied and actual rates is negative. This, however, is due to the boom-bust period during the first few years.

Table 1 presents information regarding the computed user costs and rental costs. Table 2, in turn, reports some statistics regarding housing market variables and population of the cities.

Assuming constant real prices, the user cost (as a percentage of housing prices) has been the lowest in HMA. This follows from the fact that in HMA, where the housing price level is by far the highest of the cities, M is the smallest whereas the other factors determining the computed user cost-% are the same in all the cities. Despite the lowest user cost-%, the user cost-to-rental price ratio was the greatest in HMA. This is because of the substantially higher price-to-rent ratio in HMA than in the other Finnish cities.

The average actual user costs (perfect foresight model) over 1989–2008 vary between 2.1% (in HMA) and 5.1% (in Lahti), while the average rental costs are notably larger, from 5.8% (HMA) to 9.5% (Rovaniemi). Since the user cost-% is just the inverse of the equilibrium P/E, zero real appreciation expectation justifies a gross P/E of 34 in HMA in 2008. In Rovaniemi the corresponding figure is 25. If the average long-run appreciation rates were employed as expected price growth rates, the model would rationalize even higher gross P/E ratios. In general, the larger and faster growing the city is, the smaller the user cost and rental cost percentages are.¹⁸ This is in line with the prior expectations and with the empirical

¹⁷ In 2008 the capital income tax rate was 28%, while the deduction rate for first-time buyers was 30%.

¹⁸ Table A3 in the Appendix shows OLS regression test results on the dependence of user cost % and rental cost % upon the city size and growth rate.

Table 1. *Details of annual user costs and rental costs over 1989–2008 (% of housing price level).*

	Avg. actual user cost (owner-occupants)	Avg. user cost (0% exp. real app., first-time buyers)	User cost in 2008 (0% exp. real app., first-time buyers)	Average rental cost	Rental cost in 2008
HMA	3.7%	6.0%	3.0%	5.8%	5.2%
Tampere	4.5%	6.6%	3.2%	7.5%	6.4%
Turku	6.1%	7.0%	4.1%	7.9%	7.3%
Oulu	5.5%	6.8%	3.6%	7.8%	7.6%
Lahti	6.6%	7.4%	4.5%	8.6%	8.1%
Kuopio	5.6%	6.9%	3.5%	7.7%	7.2%
Jyväskylä	5.6%	6.7%	3.5%	7.6%	7.3%
Pori	6.4%	7.6%	4.4%	8.7%	7.9%
Rovaniemi	5.3%	7.7%	4.0%	9.5%	8.2%
Vaasa	5.7%	6.9%	3.9%	7.8%	7.2%

Table 2. *Housing market and population statistics.*

	Population 1.1.2008	Avg. annual population growth 1989–2008	Avg. housing price level in 2008 (€ per square meter)	Avg. rental price level in 2008 (€ per square meter)
HMA	1,007,611	1.1%	2,949	12.7
Tampere	207,866	1.0%	1,970	10.5
Turku	175,286	0.5%	1,658	10.1
Oulu	131,585	1.5%	1,534	9.8
Lahti	99,308	0.3%	1,430	9.7
Kuopio	91,320	0.6%	1,710	10.2
Jyväskylä	85,402	1.1%	1,697	10.4
Pori	76,403	–0.0%	1,261	8.3
Rovaniemi	58,825	0.6%	1,317	9.0
Vaasa	58,597	0.4%	1,501	9.1

The average housing price level is the average transaction price of privately financed flats. The average rental price level, in turn, represents the average rent of privately financed rental dwellings. The annual population growth rate is the average annual logarithmic growth rate over 1.1.1989–1.1.2008.

result by Himmelberg et al. (2005) from the US.

The notable differences between the cities show clearly that it is reasonable to use city level data instead of national level data when analysing potential misalignments in housing prices. As Table 1 and Figures 1–5 demonstrate, the use of national level data would conceal substantial regional differences and therefore might give rise to misleading conclusions regarding a number of regional housing markets.

Correlation between the materialized user cost-% and rental cost-% over time is negative and highly significant in all the cities. This is unsurprising,

since in the beginning of the sample period price level dropped from the overly high level resulting high user costs while the high price level compared to rental prices, naturally, caused low rent-to-price ratios. Similarly, when prices started to gradually rise in the late 1990s after the recession, user costs became small whereas rental cost-% was high due to the housing price overreaction downwards during the early 1990s.

5 Summary and conclusions

The often used housing price-to-income and housing price-to-rent ratios are problematic in housing market analysis and may induce misleading conclusions. Instead, a theoretically sound basis for evaluating if housing prices are misaligned is given by the housing market no-arbitrage condition. Unfortunately, the empirical application of the no-arbitrage condition has notable complications. This article reviews these complications and previous empirical literature that has employed the condition. The main contribution of this article lies in the suggested solutions to overcome some of the complications and in the empirical analysis conducted using data from Finland. The empirical analysis includes some similar complications to those in the previous related literature, though.

The paper recommends the computation of implied expected appreciation derived from the no-arbitrage condition. By reporting the implied expected housing price growth, i.e. the appreciation rate at which user cost equals rental cost, a researcher makes it easy for a reader to adjust the results to diverging assumptions regarding factors such as the risk premium or the expected inflation rate. In the computations, the real appreciation rate is better to use than the nominal one, since during times of high inflation large implied expected nominal appreciation would not necessarily indicate high current price level.

The paper also claims that the maintenance costs as a fraction of housing prices are expected to be smaller in larger cities and, most importantly, may be downward trending in a growing metro area. Hence, to make reliable conclusions the constant maintenance cost assumption should not be used when comparing the user costs between regions or when studying the evolution of user costs over time.

Moreover, if the price-to-rent ratio is used, for simplicity reasons for instance, the rental cash flow net of maintenance costs should be employed. This is because, in general, maintenance costs do not grow at the same rate as rental prices do. In particular, in a growing metro area rental prices are expected to grow faster than maintenance costs and thereby the equilibrium price-to-rent (gross) ratio is likely to trend upwards. It is also clear that, in general, the price-to-rent ratio cannot be utilized to compare housing price “fairness” across cities.

The empirical application of the no-arbitrage relation using data from ten cities in Finland implies that housing price level has not been based on high expected appreciation in the 2000s. This attenuates fears for a notable overpricing of housing. The findings are in line with the calculations of Girouard et al. (2006) concerning the whole of Finland during 1995–2004. If housing supply has not been able to adjust properly to the growing housing demand during the last 10 years and thereby the supply will converge towards the market equilibrium with a notable

lag, the no-arbitrage condition may give a slightly misleading picture regarding the attractiveness of owner-occupied housing in 2008. The other variables being constant, a substantial increase in the housing stock would cause downward pressure in the rents and prices in the housing market. That is, the empirical model cannot totally exclude the possibility of a notable overpricing. Moreover, the current global financial crisis has deteriorated many of the fundamentals that affect housing demand. Note also that the tightening of liquidity constraints due to the crisis may cause housing price level to drop away from the no-arbitrage condition.

In any matter, the analysis shows that there has not been a housing price bubble in the sense that the prevailing price level would have been based on speculative expectations of ongoing rapid housing price increases. Instead, it appears that the rapid price growth since the mid 1990s has been, to a large extent, adjustment of housing prices towards the no-arbitrage relation. Housing prices dropped to an overly low level compared with rental prices during the deep recession in the early and mid 1990s. The notable loosening in the credit constraints during the last ten years has facilitated the ability of housing price level to move towards the no-arbitrage condition.

The analysis further proposes that the high deductibility rates of mortgage interest payments in taxation prior to 1993, together with the relatively high inflation and nominal interest rates, fortified the housing price overshoot in the late 1980s as the availability of credit increased remarkably. After 1992, in turn, the user cost of owner-occupied housing has been extremely low on average. This suggests that households have highly underestimated the future housing price growth rate since the early 1990s. Another explanation would be that the risk premium for housing is substantially greater in Finland than suggested in the literature. This explanation, however, would require unreasonably large risk premiums. One further factor affecting the divergence between user costs and rental cost may have been the credit constraints faced by households.

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Appendix

Table A1. *DF-GLS unit root test statistics on the P/Y and net P/E series over 1985–2008.*

City	P/Y	P/E
HMA	-.42 (2)	-2.58** (2)
Tampere	-.11 (2)	-1.72* (2)
Turku	-.40 (2)	-2.01** (2)
Oulu	-41 (0)	-74 (1)
Lahti	-55 (2)	-1.32 (2)
Kuopio	-55 (2)	-1.11 (2)
Jyväskylä	-.46 (1)	-.96 (1)
Pori	-.51 (2)	-1.34 (2)
Vaasa	-.94 (2)	-.88 (1)
Rovaniemi	-1.01 (0)	-1.52 (0)

Critical values at the 5% and 10% level of significance are 1.95 and 1.60, respectively. The tests do not include any deterministic variables. The number of lags included in the test is reported in the parenthesis. The lag length is based on the general-to-specific method. However, additional lags are included in case the LM(1) test clearly rejects the hypothesis of no autocorrelation in the residuals. ** denotes statistical significance at the 5% level and * at the 10% level.

Table A2. *DF-GLS unit root test statistics on the user cost-to-rent ratios employing the constant expected appreciation and perfect foresight assumptions over 1989–2008.*

City	Constant app.	Perfect foresight
HMA	-.70 (0)	-2.64** (0)
Tampere	-.73 (0)	-2.44** (0)
Turku	-.65 (0)	-2.68** (0)
Oulu	-.71 (0)	-1.77* (0)
Lahti	-.64 (0)	-2.53** (0)
Kuopio	-.76 (0)	-2.39** (0)
Jyväskylä	-.76 (0)	-2.36** (0)
Pori	-.70 (0)	-1.91* (0)
Vaasa	-.70 (0)	-3.83** (1)
Rovaniemi	1.87* (0)	-2.95** (0)

Critical values at the 5% and 10% level of significance are 1.95 and 1.62, respectively. The tests do not include any deterministic variables. The number of lags included in the test is reported in the parenthesis. The lag length is based on the general-to-specific method. However, additional lags are included in case the LM(1) test clearly rejects the hypothesis of no autocorrelation in the residuals. ** denotes statistical significance at the 5% level and * at the 10% level. The ratios are the ones regarding owner occupants.

Table A3. Ordinary Least Squares (OLS) test results regarding the dependence of user cost % and rental cost % on the city size and growth rate.

Explanatory variable	Dependent variable	
	User cost % 2008	Rental cost % 2008
Constant	0.044* (0.002)	0.080* (0.003)
City size	-0.007* (0.004)	-0.025* (0.006)
Growth rate	-0.641* (0.240)	-0.354 (0.371)
R²	0.61	0.72
JB	0.88	0.72

The user cost % is that of the first-time home-buyers using the constant real housing price level assumption. City size is the population of the city in 2008 (millions of inhabitants) and growth rate is the average annual population growth rate of the city during 1989–2008. Standard errors of the coefficient estimates are reported in the parenthesis. R² is the coefficient of determination and JB stand for the p-value in the Jarque-Bera test for the residual normality. * denotes statistical significance of the coefficient estimate at the 10% level of significance.